

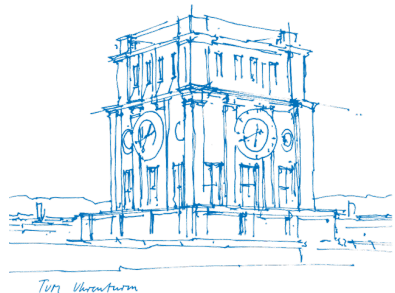
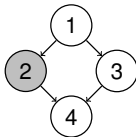
# Max-linear Bayesian networks

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## Graphical models [Lauritzen (1996)]

- Represent multivariate distributions to facilitate statistical analysis.
- Describe high-dimensional distribution by a careful combination of lower dimensional factors.
- Use graphs as natural data structure models for algorithmic treatment.
- Use graphical models for causal interpretation through a recursive system on a directed acyclic graph (DAG) [Pearl (2009)].
- Conditional independence and Markov properties are essential features.

We present **conditional independence properties** of **max-linear Bayesian networks**, which emphasize the difference to **Bayesian linear networks**.

## Max-linear Bayesian networks (MLBN) [Gissibl & K. (2018)]

Let  $\mathcal{D} = (V, E)$  be a DAG and each node  $i$  represent a r.v.  $X_i$ .

Define the **MLBN** over  $\mathcal{D}$  by the **recursive ML structural equation system** [Pearl (2009)]

$$X_i := \bigvee_{k \in \text{pa}(i)} c_{ki} X_k \vee Z_i \quad i = 1, \dots, d$$

for independent **innovations**  $Z_1, \dots, Z_d > 0$  with continuous distributions, **coefficients**  $c_{ki} > 0$ , and  $\text{pa}(i)$  (parents of  $i$ ) denotes the set of nodes  $j$  with a directed edge from  $j$  to  $i$  ( $j \rightarrow i$ ).

The system has solution

$$X_i = \bigvee_{j \in \text{an}(i) \cup \{i\}} c_{ij}^* Z_j \quad i = 1, \dots, d.$$

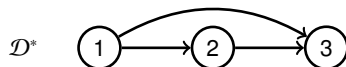
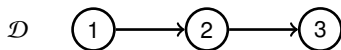
where  $\text{an}(i)$  (ancestors of  $i$ ) denotes the set of nodes  $j$  with a **directed path** from  $j$  to  $i$  ( $j \rightsquigarrow i$ ), and

$c_{ij}^*$  is a maximum taken over all the products of the coefficient along  $j \rightsquigarrow i$ . Any such path that realizes this maximum is called **max-weighted** under  $C$ .

## Path notation

- Define  $C^* = (c_{ij}^*)$  such that  $c_{ij}^*$  is a **maximum weight** of all paths (weight = product of the coefficients) along  $j \rightsquigarrow i$ . Hence,  $C^*$  is a weighted reachability matrix, i.e., supported by the **reachability DAG**  $\mathcal{D}^*$ .

**Example.**  $\mathcal{D} = (V, E) = (\{1, 2, 3\}, \{(1, 2), (2, 3)\})$



- A **path** in a DAG  $\mathcal{D}$  is a sequence of nodes  $i_0, i_1, \dots, i_k$  such that  $i_\ell \rightarrow i_{\ell+1}$  or  $i_{\ell+1} \rightarrow i_\ell$  is an edge in  $\mathcal{D}$  for each  $\ell = 0, \dots, k$ .
- A **directed path** has edges  $i_\ell \rightarrow i_{\ell+1}$  for all  $\ell$ .
- A **collider** on a path is a node  $i_\ell$  in a path such that  $i_{\ell-1} \rightarrow i_\ell \leftarrow i_{\ell+1}$ .



## Tropical linear algebra [e.g. Butkovic (2010)]

Linear Bayesian networks are based on classical linear algebra; in contrast, max-linear Bayesian networks are based on tropical linear algebra in the max-times semiring  $(\mathbb{R}_{\geq}, \odot, \cdot)$ , defined by

$$a \odot b = a \vee b = \max(a, b), \quad a \cdot b = ab \quad \text{for } a, b \in \mathbb{R}_{\geq}.$$

These operations extend to  $\mathbb{R}_{\geq}^d$  coordinate-wise and to corresponding matrix multiplication for  $R \in \mathbb{R}_{\geq}^{m \times n}$  and  $S \in \mathbb{R}_{\geq}^{n \times p}$  as

$$(R \odot S)_{ij} = \bigvee_{k=1}^n r_{ik} s_{kj}.$$

For max-linear Bayesian networks,  $X$  is Markov with respect to its DAG. **However**, the tropical linear algebra has various consequences concerning conditional independence properties and statistical analysis of the model.

## Conditional independence

Linear graphical models identify conditional independence relations through **separation criteria** applied to a graph.

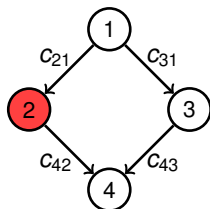
The standard separation criteria is given by the following definition.

**Definition** Two nodes  $i, j \in V$  are  **$d$ -connected** given a set  $K \in V \setminus \{i, j\}$ , if there is a path  $\pi : j \rightsquigarrow i$  such that all colliders on  $\pi$  are in  $K \cup \text{an}(K)$  and no non-collider on  $\pi$  is in  $K$ . For three disjoint subsets  $I, J, K$  of the node set  $V$ , the node set  $K$   **$d$ -separates**  $I$  and  $J$ , if no pair of nodes  $i \in I$  and  $j \in J$  is  $d$ -connected relative to  $K$ .

### Note:

- Conditional independence properties for max-linear Bayesian networks are very different from those in linear Bayesian networks. In particular, they are often not faithful to their underlying DAG  $\mathcal{D}$ .
- Hence, the above  $d$ -separation criterion on the DAG typically will not identify all valid conditional independence relations.
- This is in contrast to the situation for most Bayesian networks based on discrete random variables or linear structural equations.

## Diamond DAG: max-weighted and other paths



$1 \rightarrow 2 \rightarrow 4$  is max-weighted  $\Leftrightarrow c_{42}c_{21} \geq c_{43}c_{31}$ .  
Then

$$X_1 = Z_1, \quad X_2 = c_{21}X_1 \vee Z_2,$$

$$X_4 = c_{42}X_2 \vee Z_4 \vee c_{43}X_3$$

$$= c_{42}(Z_2 \vee c_{21}Z_1) \vee Z_4 \vee c_{43}(Z_3 \vee c_{31}Z_1)$$

$$= c_{42}Z_2 \vee c_{42}c_{21}Z_1 \vee Z_4 \vee c_{43}Z_3 \vee c_{43}c_{31}Z_1$$

$$= c_{42}Z_2 \vee c_{42}c_{21}Z_1 \vee Z_4 \vee c_{43}Z_3$$

$$= c_{42}X_2 \vee Z_4 \vee c_{43}Z_3$$

$$\Rightarrow X_1 \perp\!\!\!\perp X_4 \mid X_2.$$

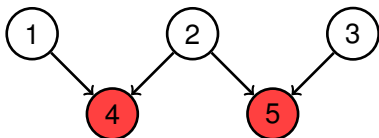
This does **not** follow from the  $d$ -separation criterion.

Here, the fact that  $1 \rightarrow 2 \rightarrow 4$  is **max-weighted** renders the path

$1 \rightarrow 3 \rightarrow 4$  unimportant for the conditional independence  $X_1 \perp\!\!\!\perp X_4 \mid X_2$ ,  
even if  $1 \rightarrow 3 \rightarrow 4$  were also max-weighted (that is, if  $c_{42}c_{21} = c_{43}c_{31}$ ).

## Cassiopeia DAG: double colliders along a path

For the sake of the argument, I set all  $c_{ij} = c_{ij}^* = 1$ .



$$X_1 = Z_1 \quad X_2 = Z_2 \quad X_3 = Z_3$$

$$X_4 = Z_1 \vee Z_2 \vee Z_4$$

$$X_5 = Z_2 \vee Z_3 \vee Z_5$$

**Indeed:**  $X_1 \perp\!\!\!\perp X_3 \mid \{X_4 = x_4, X_5 = x_5\}$  for all coefficient matrices  $C$ :

Let  $x_K = (x_4, x_5)$  and recall that all  $Z_i$  are a.s. different.

Then

$$\begin{bmatrix} x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} Z_1 \vee Z_2 \vee Z_4 \\ Z_2 \vee Z_3 \vee Z_5 \end{bmatrix} \geq \begin{bmatrix} Z_4 \\ Z_5 \end{bmatrix} \quad \text{and} \quad \begin{bmatrix} x_4 \\ x_5 \end{bmatrix} \geq \begin{bmatrix} Z_1 \vee Z_2 \\ Z_2 \vee Z_3 \end{bmatrix}$$



## Cassiopeia continued

We have three situations for  $(x_4, x_5)$  corresponding to

$$x_4 > x_5,$$

$$x_4 < x_5,$$

$$x_4 = x_5$$

$$\begin{bmatrix} x_4 \\ x_5 \end{bmatrix} \geq \begin{bmatrix} Z_1 \\ Z_2 \vee Z_3 \end{bmatrix}, \quad \begin{bmatrix} x_4 \\ x_5 \end{bmatrix} \geq \begin{bmatrix} Z_1 \vee Z_2 \\ Z_3 \end{bmatrix}, \quad \begin{bmatrix} x_4 \\ x_5 \end{bmatrix} \geq \begin{bmatrix} Z_1 \\ Z_3 \end{bmatrix} \text{ and } Z_2 = x_4 = x_5.$$

Hence, all  $Z_i$  are bounded in all three cases. Moreover,  $Z_1$  and  $Z_3$  never occur together in any inequality, rendering  $X_1 \perp\!\!\!\perp X_3 \mid X_{\{4,5\}}$ .

$x_4 > x_5$ : then the **causal source** of 4 is 1, and of 5 they are 2,3

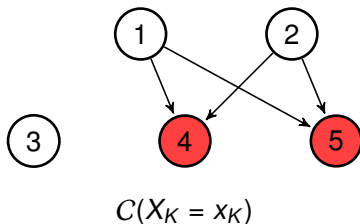
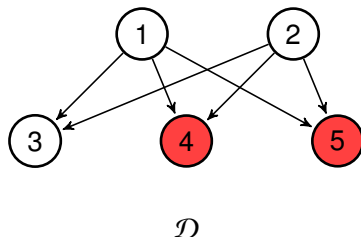
$x_4 < x_5$ , then the situation is reversed

$x_4 = x_5$ , then  $X_2 = x_4 = x_5$  is a **fixed node**

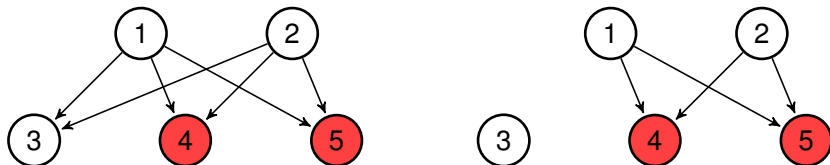
**Note:** Conditional independence does not follow from the  $d$ -separation criterion, since the path from 1 to 3 is  $d$ -connecting relative to  $\{4, 5\}$ .

## Tent DAG: context specific conditional independence

The **source graph**  $C(X_K = x_K)$ :



**Figure:** Left: Tent DAG  $\mathcal{D}$ . Right: For all coefficients equal to 1, the source graph  $C(X_K = x_K)$  with observed values  $x_4 = x_5 = 2$  is obtained from  $\mathcal{D}$  by removing the edges  $1 \rightarrow 3$  and  $2 \rightarrow 3$ , which become redundant in the context  $\{X_4 = X_5 = 2\}$ .

Continuing the Tent DAG with  $\{X_4 = X_5 = 2\}$ 

$$X_1 = Z_1, \quad X_2 = Z_2, \quad X_3 = Z_3 \vee X_1 \vee X_2$$

$$X_4 = Z_4 \vee X_1 \vee X_2 = 2, \quad X_5 = Z_5 \vee X_1 \vee X_2 = 2.$$

Since  $Z_1, \dots, Z_5$  are a.s. different, it holds outside a null-set that  $X_1 \vee X_2 = Z_1 \vee Z_2 = 2$ . This introduces bounds on the innovations; we must have  $Z_1, Z_2, Z_4, Z_5 \leq 2$  and it also holds that  $X_3 \geq 2$ . Further, we then have

$$X_1 = Z_1, \quad X_2 = Z_2, \quad X_1 \vee X_2 = 2, \quad X_3 = Z_3 \vee 2,$$

$$X_4 = Z_4 \vee 2 = 2, \quad X_5 = Z_5 \vee 2 = 2,$$

Now, the dependence of  $X_3$  on  $X_1, X_2$  has disappeared, hence  $X_3 \perp\!\!\!\perp (X_1, X_2) \mid X_4 = X_5 = 2$ . This independence statement is reflected in the lack of edges  $1 \rightarrow 3$  and  $2 \rightarrow 3$  in the source DAG  $C(X_4 = X_5 = 2)$ .

## Summary and conclusion

- MLBN models have very different conditional independence properties than LBN models.
- I have given 3 examples, where the classical  $d$ -separation criterion fails.
- For an observed set of nodes  $K$ , a representation of  $X_{\bar{K}} \mid X_K = x_K$  guides us to find a **reduced representation of  $X_{\bar{K}}$**  taking deterministic features of a MLBN into account.
- **Impact graphs** describe **how extreme events spread** in the MLBN.
- The union of all **impact graphs compatible with  $\{X_K = x_K\}$**  is the starting point for tracking **possible sources** of  $\{X_K = x_K\}$ .
- Cleaning up this union of graphs for fixed and redundant nodes and redundant edges yields the **source graph  $C(X_K = x_K)$**  giving a compact representation of the conditional distribution given  $X_K = x_K$ .
- A new **\*-separation criterion** is equivalent to CI statements in context-free and context-dependent settings, which we formulate as \*-separation in different derived DAGs.

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